Exchange Rate Regime Analysis for the Chinese Yuan

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Abstract

We investigate the Chinese exchange rate regime after China gave up on a fixed exchange rate to the US dollar in 2005. This reproduces the analysis from Zeileis, Shah, and Patnaik (2010) initiated by Shah, Zeileis, and Patnaik (2005). Please refer to these papers for a more detailed discussion.

1 Analysis

Exchange rate regime analysis is based on a linear regression model for cross-currency returns. A large data set derived from exchange rates available online from the US Federal Reserve at http://www.federalreserve.gov/releases/h10/Hist/ is provided in the FXRatesCHF data set in fxregime.

> library("fxregime")
> data("FXRatesCHF", package = "fxregime")

It is a "zoo" series containing 25 daily time series from 1971-01-04 to 2010-02-12. The columns correspond to the prices for various currencies (in ISO 4217 format) with respect to CHF as the unit currency.

In the following, we investigate the exchange rate regime for the Chinese yuan CNY which was fixed to the US dollar USD in the years leading up to mid-2005. In July 2005, China announced a small appreciation of CNY, and, in addition, a reform of the exchange rate regime. The People's Bank of China (PBC) announced this reform to involve a shift away from the fixed exchange rate to a basket of currencies with greater flexibility. In August 2005, PBC also announced that USD, JPY, EUR and KRW would be the currencies in this basket. Further currencies announced to be of interest are GBP, MYR, SGD, RUB, AUD, THB and CAD.

Despite the announcements of the PBC, little evidence could be found for China moving away from a USD peg in the months after July 2005 (Shah *et al.*, 2005). To begin our investigation here, we follow up on our own analysis from autumn 2005: Using daily returns for the first three months after the announcement, we establish a stable exchange regression and monitor it in the subsequent months. The currencies considered by Zeileis *et al.* (2010) are a basket of the most important floating currencies (USD, JPY, EUR, GBP). The returns can be extracted from FXRatesCHF and pre-processed via

```
> cny <- fxreturns("CNY", frequency = "daily",
+ start = as.Date("2005-07-25"), end = as.Date("2009-07-31"),
+ other = c("USD", "JPY", "EUR", "GBP"), data = FXRatesCHF)
```

In a first step, we fit the exchange regression for these first three months after the announcements of the PBC.

```
> cny_lm <- fxlm(CNY ~ USD + JPY + EUR + GBP,
+
    data = window(cny, end = as.Date("2005-10-31")))
> summary(cny_lm)
Call:
fxlm(formula = CNY ~ USD + JPY + EUR + GBP, data = window(cny,
    end = as.Date("2005-10-31")))
Residuals:
                      Median
     Min
                10
                                    30
                                             Max
-0.065697 -0.021036 0.001147 0.021440 0.069985
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.004782 0.003688 -1.297
                                           0.199
USD
             0.999653 0.008779 113.868
                                          <2e-16 ***
JPY
             0.004668
                       0.010669
                                 0.437
                                           0.663
EUR
            -0.014238 0.026516 -0.537
                                           0.593
GBP
            -0.007744
                       0.014568 -0.532
                                           0.597
___
Signif. codes:
               0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.02953 on 63 degrees of freedom
Multiple R-squared: 0.9979,
                                   Adjusted R-squared: 0.9978
F-statistic: 7577 on 4 and 63 DF, p-value: < 2.2e-16
```

Only the USD coefficient differing significantly from 0 (but not significantly from 1), thus signalling a very clear USD peg. The R^2 of the regression is 99.8% due to the extremely low standard deviation of $\sigma = 0.028$. (Note that we use the un-adjusted estimate of σ , rather than the adjusted version reported in the summary() above.)

To capture the fluctuation in the parameters during this history period, we compute the associated empirical fluctuation process

> cny_efp <- gefp(cny_lm, fit = NULL)</pre>

that can be visualized (along with the boundaries for the double maximum test) by

> plot(cny_efp, aggregate = FALSE, ylim = c(-1.85, 1.85))

Figure 1 shows that the fluctuation in the parameters during this history period is very small and non-significant:

> sctest(cny_efp)



M-fluctuation test

Figure 1: Historical fluctuation process for CNY exchange rate regime.

```
M-fluctuation test
```

```
data: cny_efp
f(efp) = 1.0968, p-value = 0.6965
```

The same fluctuation process can be continued in the monitoring period to check whether future observations still conform with the established model. Using a linear boundary, derived at 5% significance level (for potentially monitoring up to T = 4), this can be performed via

```
> cny_mon <- fxmonitor(CNY ~ USD + JPY + EUR + GBP,
+ data = window(cny, end = as.Date("2006-05-31")),
+ start = as.Date("2005-11-01"), end = 4)
> plot(cny_mon, aggregate = FALSE)
```

yielding the visualization in Figure 2. In the first months, up to spring 2006, there is still moderate fluctuation in all processes signalling no departure from the previously established USD peg. In fact, the only larger deviation during that time period is surprisingly a *decrease* in the variance—corresponding to a somewhat tighter USD peg—which almost leads to a boundary crossing in January 2006. However, the situation relaxes a bit before in the next weeks before in March 2006 the variance component of the fluctuation process starts to deviate clearly from its mean. However, none of the coefficients deviates from its zero mean, signalling that there was no significant change in the currency weights. The change occurs in

> cny_mon

Monitoring of FX model

Formula: CNY ~ USD + JPY + EUR + GBP History period: 2005-07-26 to 2005-10-31 Break detected: 2006-03-27

To capture the changes in the China's exchange rate regime more formally, we fit a segmented exchange rate regression based on the full extended data set:

```
> cny_reg <- fxregimes(CNY ~ USD + JPY + EUR + GBP,
+ data = cny, h = 20, breaks = 10)
```

[1] TRUE

We determine the optimal breakpoints for $1, \ldots, 10$ breaks with a minimal segment size of 20 observations and compute the associated segmented negative log-likelihood (NLL) and LWZ criterion. Both can be visualized via

> plot(cny_reg)



Monitoring of FX model

Figure 2: Monitoring fluctuation process for CNY exchange rate regime.

NLL decreases with every additional break but with a marked decrease only for going from 0 to 1 break. This is also reflected in the LWZ criterion that assumes its minimum for 3 break so that we choose a 3-break (or 4-segment) model. The estimated breakpoint is 2006-03-14, i.e., shortly before the boundary crossing in the monitoring procedure, confirming the findings above. The confidence interval for the break can be obtained by

```
> confint(cny_reg, level = 0.9)
```

Confidence intervals for breakpoints of optimal 4-segment partition:

LWZ and Negative Log-Likelihood



Figure 3: Negative log-likelihood and LWZ information criterion for CNY exchange rate regimes.

Call: confint.fxregimes(object = cny_reg, level = 0.9) Breakpoints at observation number: 5 % breakpoints 95 % 1 143 158 159 2 762 778 779 3 865 866 880 Corresponding to breakdates: 5 % breakpoints 95 % 1 2006-02-21 2006-03-14 2006-03-15 2 2008-07-31 2008-08-22 2008-08-25 3 2008-12-30 2008-12-31 2009-01-22

showing that the end of the low variance period can be determined more precisely than the start of the high variance period. The parameter estimates for both segments can be obtained by

> coef(cny_reg)

(Intercept)USDJPYEUR2005-07-26--2006-03-14-0.0050329730.99940960.005184123-0.0152439812006-03-15--2008-08-22-0.0249927730.9693984-0.0093215880.0255942922008-08-25--2008-12-31-0.0147701021.0307442-0.0264792090.0488530472009-01-02--2009-07-310.0013514040.98093890.008205345-0.007683415GBP(Variance)2005-07-26--2006-03-140.0068385120.00078168222006-03-15--2008-08-22-0.0128676500.01128562742008-08-25--2008-12-310.0071871780.06939695762009-01-02--2009-07-310.0085673360.0019749197

A complete summary can be computed by first re-fitting the model on both sub-samples (returning a list of "fxlm" objects) and then applying the usual summary():

```
> cny_rf <- refit(cny_reg)</pre>
> lapply(cny_rf, summary)
$'2005-07-26--2006-03-14'
Call:
fxlm(formula = object$formula, data = window(object$data, start = sbp[i],
    end = ebp[i]))
Residuals:
     Min
                1Q
                      Median
                                    ЗQ
                                             Max
-0.106628 -0.015830 0.001518 0.016454 0.090368
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.005033 0.002266 -2.221 0.0278 *
USD
            0.999410 0.005421 184.370 <2e-16 ***
JPY
            0.005184 0.005230 0.991
                                          0.3231
EUR
           -0.015244 0.016588 -0.919
                                          0.3596
            0.006839 0.008257 0.828
                                          0.4088
GBP
___
               0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Signif. codes:
Residual standard error: 0.02841 on 153 degrees of freedom
Multiple R-squared: 0.9979,
                                   Adjusted R-squared: 0.9978
F-statistic: 1.788e+04 on 4 and 153 DF, p-value: < 2.2e-16
$'2006-03-15--2008-08-22'
Call:
fxlm(formula = object$formula, data = window(object$data, start = sbp[i],
    end = ebp[i]))
```

Residuals: Min 1Q Median ЗQ Max -0.44652 -0.06071 0.01135 0.06138 0.45665 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) -0.024993 0.004300 -5.812 9.92e-09 *** 0.969398 0.011533 84.054 < 2e-16 *** USD JPY -0.009322 0.010450 -0.892 0.373 0.025594 0.022943 1.116 EUR 0.265 -0.012868 0.012147 -1.059 0.290 GBP ___ Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 0.1067 on 615 degrees of freedom Multiple R-squared: 0.9649, Adjusted R-squared: 0.9646 F-statistic: 4223 on 4 and 615 DF, p-value: < 2.2e-16 \$'2008-08-25--2008-12-31' Call: fxlm(formula = object\$formula, data = window(object\$data, start = sbp[i], end = ebp[i])) Residuals: Min 1Q Median 3Q Max -0.97806 -0.11418 -0.01290 0.09812 0.87997 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) -0.014770 0.029756 -0.496 0.621 USD 1.030744 0.043672 23.602 <2e-16 *** JPY $-0.026479 \quad 0.030149 \quad -0.878 \quad 0.382$ 0.048853 0.058852 0.830 0.409 EUR GBP 0.007187 0.035289 0.204 0.839 ___ Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 0.2713 on 83 degrees of freedom Multiple R-squared: 0.9562, Adjusted R-squared: 0.954 F-statistic: 452.6 on 4 and 83 DF, p-value: < 2.2e-16

\$'2009-01-02--2009-07-31'

```
Call:
fxlm(formula = object$formula, data = window(object$data, start = sbp[i],
    end = ebp[i]))
Residuals:
                 1Q
                        Median
      Min
                                      ЗQ
                                                Max
-0.225789 -0.021239 -0.000453
                               0.019380
                                          0.145003
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept)
             0.001351
                         0.003734
                                    0.362
                                             0.7179
USD
             0.980939
                         0.005081 193.060
                                             <2e-16 ***
JPY
                         0.004325
             0.008205
                                    1.897
                                             0.0598 .
FUR.
            -0.007683
                         0.009480
                                   -0.810
                                             0.4190
GBP
             0.008567
                         0.004464
                                    1.919
                                             0.0570 .
___
                0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Signif. codes:
Residual standard error: 0.04521 on 143 degrees of freedom
Multiple R-squared: 0.9979,
                                     Adjusted R-squared:
                                                           0.9978
F-statistic: 1.699e+04 on 4 and 143 DF, p-value: < 2.2e-16
```

These results allow for several conclusions about the Chinese exchange rate regime after spring 2006: CNY was still closely linked to USD. The exchange rate regime got much more flexible increasing from $\sigma = 0.028$ to 0.106 which is still very low, even compared with other pegged exchange rate regimes (see the results India in vignette("INR", package = "fxregime")). The intercept was clearly smaller than 0, reflecting a slow appreciation of the CNY and thus signalling a modest liberation of the rigid USD peg in spring 2006. Towards the end of 2008, the modest liberation was abandoned again and since 2009 the exchange rate regime is again an extremely tight USD peg without appreciation.

2 Summary

For the Chinese yuan, a 4-segment model is found for the time after July 2005 when China gave up on a fixed exchange rate to the USD. While being closely linked to USD in all periods, there had been small steps in the direction of the claims of the Chinese central bank: flexibility slightly increased while the weight of the USD in the currency basket slightly decreased. However, these steps were reversed again towards the end of 2008.

References

Shah A, Zeileis A, Patnaik I (2005). "What is the New Chinese Currency Regime?" *Report 23*, Department of Statistics and Mathematics, WU Wirtschaftsuniversität Wien, Research Report Series. URL http://epub.wu.ac.at/.

Zeileis A, Shah A, Patnaik I (2010). "Testing, Monitoring, and Dating Structural Changes in Testing, Monitoring, and Dating Structural Changes in Exchange Rate Regimes." Computational Statistics & Data Analysis, 54(6), 1696–1706. doi:10.1016/j.csda.2009.12.005.